The effect of the COVID-19 health disruptions on breast cancer risk: A semi-Markov modelling approach

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Outline

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- A projection model for breast cancer
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Motivation

Breast cancer (BC) is

- the most common cancer diagnosed in women
- one of the leading causes of death for women

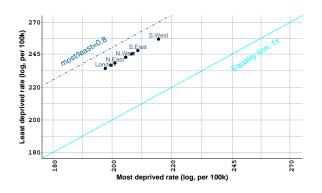
Investigate BC rates in the presence of:

 major disruptions to health services, particularly caused by a catastrophic event, e.g. the COVID-19, preventing or delaying the diagnosis of BC

Projection of BC mortality into the future



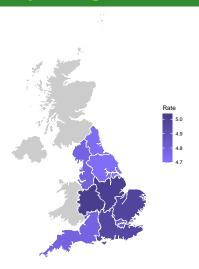
Most v. least deprived by region: BC incidence in England - 2017



- Not a life-style cancer
- Rates for least deprived higher (higher screening?)
- Less regional variation as compared to, e.g., lung cancer



Regional variation: BC mortality in England - 2018



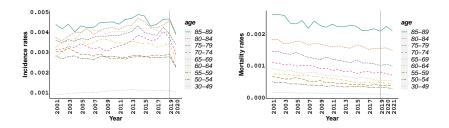
✓ Rate is per 10K✓ Deprivation is not significant



What insights we gain from BC data

- Socio-economic differences are less relevant as compared to, e.g., lung cancer incidence/mortality
- Not (easily) controllable or preventable risk factors
- Regional inequality exists but relatively low
 - High BC screening awareness
 - National BC screening programme for ages 47–73
- The availability of BC screening is crucial for early diagnosis, as BC can be curable

BC incidence and mortality in England: COVID years



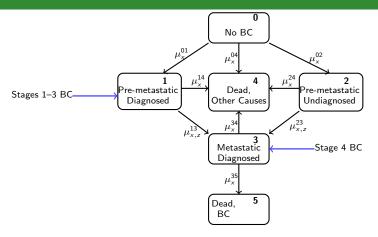
 A significant decline in BC incidence, as low as 25% at ages 60–64, in 2020 as compared to the same period in 2019

Incidence (left) v. Mortality (right)

• An increase in BC mortality from ages 65+, as high as 7%, in 2020 as compared to the same period in 2019



Multi-state model for BC transitions



- 'Dead from BC' is only accessible from 'Metastatic Diagnosed'
- Onset of BC remains unchanged $\Rightarrow \mu_{\rm x}^{01} + \mu_{\rm x}^{02} = \mu_{\rm x}^*$
- Treatment is available in 'Pre-metastatic Diagnosed' NOT in 'Pre-metastatic Undiagnosed' $\Rightarrow \mu_{x,z}^{13} < \mu_{x,z}^{23}$

A convenient parametrisation of the model

From

$$\mu_{\rm x}^{\rm 01} + \mu_{\rm x}^{\rm 02} = \mu_{\rm x}^*$$

we can write

$$\begin{split} \mu_x^{01} &= \alpha \, \mu_x^* \\ \mu_x^{02} &= \left(1 - \alpha\right) \mu_x^*, \qquad 0 < \alpha < 1 \end{split}$$

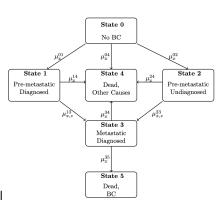
 α : level of BC diagnoses

Also we assume

$$\mu_{x,z}^{13} = \beta \, \mu_{x,z}^{23}, \qquad \beta < 1$$

 β : availability of BC treatment Transitions to death due to other causes from all 'live' states are equal to $\mu_{\rm x}^{\rm 04}$

$$\mu_{x}^{14} = \mu_{x}^{24} = \mu_{x}^{34} = \mu_{x}^{04}$$



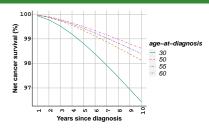
BC semi-Markov model: pre-Covid rates

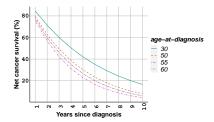
| Age | $\mu_{\scriptscriptstyle X}^{01}$ | $\mu_{\scriptscriptstyle X}^{04}$ | $\mu_{\scriptscriptstyle X}^{35}$ |
|-------|-----------------------------------|-----------------------------------|-----------------------------------|
| 30–49 | 0.00086 | 0.00084 | 0.16739 |
| 50-54 | 0.00224 | 0.00228 | 0.24005 |
| 55-59 | 0.00233 | 0.00363 | 0.24005 |
| 60-64 | 0.00282 | 0.00588 | 0.28060 |
| 65-69 | 0.00318 | 0.00952 | 0.28060 |
| 70-74 | 0.00280 | 0.01643 | 0.36002 |
| 75-79 | 0.00311 | 0.02987 | 0.40000 |
| 80-84 | 0.00338 | 0.05496 | 0.49711 |
| 85–89 | 0.00362 | 0.10112 | 0.50000 |

- $\mu_{\rm x}^{01}$: ONS/NHS Digital data, 81% of new BC registrations, England, 2001–2019
- \bullet μ_{x}^{04} : ONS data, deaths from other causes, England, 2001–2019
- $\mu_{x,z}^{13}$: Average metastasis rates per 1000 person-years (Colzani et al., 2014)
- μ_x^{35} : BC deaths by age within 12 months after Stage 4 BC diagnosis (Zhao et al., 2020)



BC net survival: pre-Covid rates





Pre-metastatic BC (left) v. Metastatic BC (right)

- ullet Baseline scenarios are carried out for women when lpha= 0.6 and $eta=rac{1}{7}$
- Net Survival: ONLY consider 'Dead, BC' as cause of death AFTER BC diagnosis
- An unusual age pattern in pre-metastatic BC net survival
- Lower metastatic BC net survival at older ages

For a woman aged x, diagnosed with pre-metastatic BC, BC survival in t years:

$$\frac{1 - {}_t p_{\scriptscriptstyle X}^{14} - {}_t p_{\scriptscriptstyle X}^{15}}{1 - {}_t p_{\scriptscriptstyle x}^{14}}$$



BC semi-Markov model - COVID scenario

In order to quantify the impact of COVID-19 on BC mortality at older ages, we have

- Excess deaths from other causes, i.e. increase in $\mu_{\rm x}^{04}$
- Decline in BC diagnosis, i.e. slowdown in $\mu_{\rm x}^{01}$ and increase in $\mu_{\rm x}^{02}$

| Pandemic period | $\mu_{_{X}}^{01}/\mu_{_{X}}^{02}$ | μ | 04 × |
|---------------------|-----------------------------------|-------|---------|
| | α | 65–84 | 85–89 |
| April-Nov. 2020 | 0.8 | 1.13 | 1.12 |
| Dec. 2020-Nov. 2021 | 1 | 1.13 | 1.12 |
| Dec. 2021-Dec. 2022 | 1 | 1.10 | 1.09 |
| JanDec. 2023 | 1 | 1.07 | 1.06 |
| JanDec. 2024 | 1 | 1.04 | 1.03 |

Short-term implications up to 5 years

| Occupancy Probabilities | | | | | | | | | | |
|--------------------------|-----------------------|--------------------------|------------------|-----------------------|------------------|------------------|-----------------------|---------------------|---------------------|---------------------|
| | From State 0 | | | | From | State 1 | From S | State 3 | | |
| Age | $_{5}p_{\times}^{00}$ | $_{5}p_{_{\times}}^{01}$ | $_{5}p_{x}^{02}$ | $_{5}p_{\times}^{03}$ | $_{5}p_{x}^{04}$ | $_{5}p_{x}^{05}$ | $_{1}p_{\times}^{15}$ | $_{5}p_{_{X}}^{15}$ | $_{1}p_{_{X}}^{35}$ | $_{5}p_{_{X}}^{35}$ |
| | (%) | (%) | (%) | (%) | (%) | (%) | (%) | (%) | (%) | (%) |
| Pre-pandemic calibration | | | | | | | | | | |
| 65-69 | 93.09 | 1.47 | 0.68 | 0.31 | 4.29 | 0.16 | 0.16 | 5.98 | 24.36 | 74.15 |
| 70-74 | 90.49 | 1.22 | 0.57 | 0.23 | 7.32 | 0.16 | 0.20 | 6.82 | 30.02 | 81.25 |
| 75-79 | 85.07 | 1.31 | 0.61 | 0.24 | 12.59 | 0.19 | 0.22 | 6.97 | 32.56 | 82.61 |
| 80-84 | 75.07 | 1.26 | 0.59 | 0.20 | 22.66 | 0.21 | 0.26 | 7.21 | 38.26 | 84.79 |
| 85-89 | 59.71 | 1.07 | 0.50 | 0.17 | 38.36 | 0.19 | 0.25 | 6.29 | 37.65 | 79.54 |
| Pandemic scenario | | | | | | | | | | |
| 65-69 | 92.73 | 1.42 | 0.70 | 0.32 | 4.66 | 0.17 | 0.16 | 5.96 | 24.36 | 74.03 |
| 70-74 | 89.90 | 1.18 | 0.58 | 0.24 | 7.93 | 0.17 | 0.20 | 6.79 | 30.00 | 81.03 |
| 75-79 | 84.09 | 1.25 | 0.62 | 0.24 | 13.60 | 0.20 | 0.22 | 6.91 | 32.53 | 82.24 |
| 80-84 | 73.42 | 1.20 | 0.59 | 0.21 | 24.36 | 0.22 | 0.26 | 7.10 | 38.20 | 84.15 |
| 85–89 | 57.53 | 1.00 | 0.49 | 0.17 | 40.61 | 0.20 | 0.25 | 6.12 | 37.55 | 78.56 |

- 3–6% decline in age-specific, $_5p_{_X}^{01}$, 'Pre-metastatic Diagnosed'
- 3–5% increase in, $_5p_{_X}^{03}$, 'Metastatic Diagnosed' (Vulnerability? Higher deaths from BC and other causes?)



Changes in BC pre- v. post-pandemic

| | Addition | al deaths | Y | LL | AC in | BC mort | tality from (%) | | |
|----------------------------|----------|-----------|---------|---------|----------------|---------|-----------------|---------|--|
| | Dead | Dead | Dead | Dead | Pre-metastatic | | Metastatic | | |
| | (Other) | (BC) | (Other) | (BC) | Diagnosed | | | | |
| | State 4 | State 5 | State 4 | State 5 | State 1 | | Sta | State 3 | |
| Pre-pandemic calibration v | r | | | | 1 | E | 1 | E | |
| Pandemic scenario | | | | | 1 year | 5 year | 1 year | 5 year | |
| 65–6 | 9 363 | 10 | 7010 | 193 | 0.00 | -0.02 | 0.00 | -0.12 | |
| 70–7 | 4 607 | 9 | 9293 | 138 | 0.00 | -0.03 | -0.02 | -0.22 | |
| 75–7 | 9 1012 | 10 | 11770 | 116 | 0.00 | -0.06 | -0.03 | -0.37 | |
| 80–8 | 4 1699 | 9 | 14340 | 76 | 0.00 | -0.11 | -0.06 | -0.64 | |
| 85-8 | 9 2253 | 6 | 13158 | 35 | 0.00 | -0.17 | -0.10 | -0.98 | |

- 100,000 women in each age group, in 'No BC' at time zero, taken as January 1, 2020
- 3-6% increase in 'Dead from BC' and 5-8% increase in 'Dead from Other Causes' for women, with 'No BC' at time zero, across different ages over 5 years
- Absolute change (AC) in BC mortality is less than 1%

Years of life expectancy lost (YLL) from a given cause is:

$$YLL_{x,t}^{cause} = D_{x,t}^{cause} e_x$$

where $D_{x,t}^{\text{cause}}$ is age- and type-specific additional deaths; and

 e_x is defined using standard life tables



Sensitivity analysis

- Sensitivity analysis is carried out, all else equal, with
 - $\alpha =$ 0.4 and $\alpha =$ 0.8 (lower v. higher BC diagnoses)
 - $\beta = \frac{1}{5}$ and $\beta = \frac{1}{10}$ (worse v. better BC treatment)
 - $\mu_{\rm x}^{35}$ is 20% lower and higher than the pre-pandemic level (lower v. higher BC deaths)
- Consistent results in relation to relative changes in BC mortality and deaths from different causes, under pre- and post-pandemic scenarios

A projection model: BC mortality in England, 2001–2018

$$\begin{split} D_{a,t,r} &\sim \mathsf{Poisson}(\theta_{a,t,r} \ E_{a,t,r}) \\ \theta_{a,t,r} &\sim \mathsf{Lognormal}(\mu_{a,t,r}, \sigma^2) \\ \mu_{a,t,r} &= \beta_0 + \beta_{1,a} + \beta_{2,t} + \beta_{3,r} \\ \sigma^2 &\sim \mathsf{Inv.Gamma}(1,0.1) \\ \beta_0, \ \beta_1, \ \beta_3 &\sim \mathsf{Normal}(0,10^4) \quad [\mathsf{vague priors for risk factor effects}] \end{split}$$

Add random walk with drift for 'period' effect:

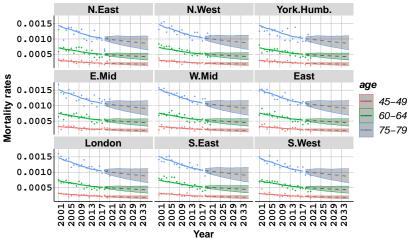
$$\begin{split} \beta_{2,t} &= \mathsf{drift} + \beta_{2,t-1} + \epsilon_t \\ \mathsf{drift} &\sim \mathsf{Normal}(0, \sigma_{\mathsf{drift}}^2) \\ \epsilon_t &\sim \mathsf{Normal}(0, \sigma_{\beta_2}^2) \\ \sigma_{\beta_2}^2 &\sim \mathsf{Inv.Gamma}(1, 0.001) \end{split}$$

for
$$t=2001,2002,\ldots,2018$$
, where $\hat{\sigma}_{\mathsf{drift}}^2=\frac{\hat{\sigma}_{\beta_2}^2}{2018-2001}$

- $D_{a,t,r}$: number of cancer deaths at age a in year t for region r
- E_{a,t,r}: mid-year population estimates



Projected mortality: BC mortality - women, ages 45–79, 2001–2036



- Decreasing trend over time
- Projected rates for youngest & oldest screening age groups NOT overlapping

Summary

- More equality in BC as compared to life-style cancers
- A valuable model relating to delays in the provision of BC diagnostic and treatment services
- As compared to the pre-pandemic scenario
 - 3-6% increase in deaths from BC and 5-8% in other causes between ages 65-89
 - Less than a 1% change in the probability of death for women with pre-metastatic BC (sp_x^{15})
 - A relatively significant change in the probability of death for women with metastatic BC $(5\rho_x^{35})$ as compared to women with pre-metastatic BC
- Less than 1% change in net single premiums when key transition rates are defined including COVID years
- Projection for BC mortality shows persistent age gap
- Duration dependence matters in actuarial applications
- Measuring parameter and model uncertainty?



More details in:

- Arık, A., Cairns, A., Dodd, E., Macdonald, A.S., Shao, A., Streftaris, G. Insurance pricing for breast cancer under different multiple state models, working paper.
- Arık, A., Cairns, A., Dodd, E., Macdonald, A.S., Streftaris, G. The effect of the COVID-19 health disruptions on breast cancer mortality for older women: A semi-Markov modelling approach, https://arxiv.org/abs/2303.16573.
- Arık, A., Cairns, A., Dodd, E., Macdonald, A.S., Streftaris, G. Estimating the impact of the COVID-19 pandemic on breast cancer deaths among older women, Living to 100 Research Symposium, 16 February 2023, conference monograph.
- Arık, A., Dodd, E., Cairns, A., Streftaris, G. Socioeconomic disparities in cancer incidence and mortality in England and the impact of age-at-diagnosis on cancer mortality, PLOS ONE, 2021.
- Arık, A., Dodd, E., Streftaris, G. Cancer morbidity trends and regional differences in England - a Bayesian Analysis, PLOS ONE, 2020.



Thank You!

Questions?

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